

UNITED STATES DEPARTMENT OF COMMERCE Economics and Statistics Administration U.S. Census Bureau Washington, DC 20233-0001



MAY 1 8 2015

MEMORANDUM FOR

Jason Fields

Survey Director, Survey of Income and Program Participation

Associate Director for Demographic Programs

From:

James B. Treat James B. Freat

Chief, Demographic Statistical Methods Division

Subject:

Nonresponse Bias Analysis for Waves 3-16 of the 2008 Survey of

Income and Program Participation (SIPP) (ALYS-15)

This document describes the results of the nonresponse bias analysis for Waves 3-16 of the 2008 Survey of Income and Program Participation (SIPP).

If you have questions, please contact Tracy Mattingly at 301-763-6445 or via email at Tracy.L.Mattingly@census.gov or Jamie Choi at 301-763-4554 or via email at Jamie.Choi@census.gov.

cc:

D. Doyle

(ADDP)

M. Marlay

T. Velkoff

(SEHSD)

J. Farber

(DSMD)

J. Scott

T. Mattingly

J. Choi

R. Culver III

M. Herbstritt

S. Mack

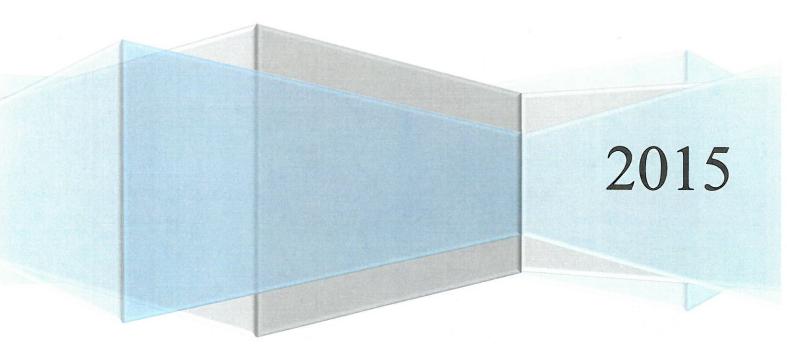
M. Sundukchi

A. Westra

U.S. Census Bureau

Nonresponse Bias Analysis for Waves 3-16 of the 2008 Survey of Income and Program Participation (SIPP)

By Jamie Choi and Tracy Mattingly



Any views expressed are those of the authors and not necessarily those of the U.S. Census Bureau.

Introduction

For a household survey, unit nonresponse is the failure to obtain any survey measures on a sample unit (e.g. household). These nonresponse rates have been increasing in recent years, even for the large government surveys. With this increasing nonresponse there is also growing concern over data quality and losing valuable information from these nonrespondents. Declining response rates can be an indicator of nonresponse bias, or a difference in survey measures between respondents and nonrespondents, which can affect data quality. However, there is not always a direct link between response rates and nonresponse bias since nonresponse bias can vary across different statistics in the same survey; low response rates in surveys may yield to some statistics having large nonresponse bias [3]. Therefore, understanding and measuring nonresponse bias for key estimates is an important aspect in determining overall data quality.

Because policy makers use estimates from the demographic surveys conducted by the U.S. Census Bureau and other agencies to determine the successfulness of programs or for national economic indicators, the highest data quality is necessary. This has led more surveys to investigate nonresponse bias, especially after the Office of Management and Budget (OMB) released their standards in 2006 that require survey programs to make plans for a nonresponse bias analysis if unit response rates fall below 80% [8]. The Census Bureau incorporated this guideline into its own standards along with guidelines which state that serious data quality issues related to nonsampling error can occur when cumulative response rates for a longitudinal survey fall below 60% and when sample attrition from wave to wave is greater than 5% [16].

The 2008 Survey of Income and Program Participation (SIPP) obtained a response rate of 80.8% at the first interview. The panel ended in December 2013 with a cumulative response rate of 49.4%.

This report presents an analysis of nonresponse bias for Waves 3-16 of the 2008 SIPP. The analysis takes into account the longitudinal aspect of the survey and utilizes the data available for those who drop out after Wave 1 to give a better picture of the overall bias associated with nonresponse. Two different studies are used to examine potential nonresponse bias. The methods include comparing key estimates of the full sample to the respondent sample and estimating the representivity of the survey using R-indicators.

Previous Research

Previous efforts to examine the bias and determine the impact on SIPP estimates included comparing characteristics of households that responded in all waves versus those that dropped out. The SIPP nonresponse workgroup found that households that were renting, living in large urban areas, and had young adults (15-24) as the householder were more likely to be nonrespondents [13]. Another study compared SIPP annual poverty rates and health insurance coverage to CPS-ASEC. Results showed differences in poverty rates at the 150% and 200% poverty thresholds and differences in health insurance coverage, especially among blacks and Hispanics [10]. Mack and Petroni [4] summarized the results from several studies that looked

into using logistic regression and various raking methodologies for nonresponse weighting adjustments. The results indicated that alternative weighting procedures did not reduce the bias more than the current adjustment for estimates of income, unemployment, government assistance and poverty.

A study completed on Wave 1 of the 2008 Panel showed through benchmark analysis that the SIPP is underestimating participation in Supplemental Nutrition and Assistance Program (SNAP), Supplemental Security Income (SSI), Medicaid, and Medicare compared to the program sources. Examining response rates across different subgroups found that those living in the Northeast had the lowest response rate out of all regions, Black householders had a lower response rate compared to Nonblack householders, and metropolitan statistical area (MSA) and principal city residents had lower response rates compared to those not living in the MSA or in the principal city. Comparing estimates of the frame variables for the full sample and for respondents-only found differences for region, urban/rural status, CBSA type, and race. The study also showed through odds ratios obtained from a logistic regression analysis that region, household size, and age of householder significantly affected the likelihood of responding [5].

Analysis completed on Wave 2 of the 2008 Panel showed by comparing estimates of the full sample and respondent sample that Wave 2 of the SIPP may be overestimating household earnings, household income, and participation in Medicare and Social security. Participation in Medicaid and SNAP may be underestimated in the responding sample. However, none of the relative differences for these key measures was greater than 5%. In addition, the large estimate of the R-indicator suggests that there is a high likelihood the respondent sample is representative of the sample or population [6].

Data

The SIPP is a longitudinal survey that collects detailed information about income, employment, health insurance, and program participation for the civilian, noninstitutionalized population living in the United States. The 2008 SIPP panel started on September 2008 and ended on December 2013 with interviews being conducted at four-month intervals called waves.

The Census Bureau employed a two-stage sample design to select the 2008 SIPP sample. A systematic selection was used to select housing units within 351 primary sampling units (PSUs) from the master address file created from the 2000 Decennial Census. In addition, households located in areas with a higher concentration of low-income households were oversampled by 44 percent to increase the accuracy of estimates for statistics of low-income households and program participation [14].

When a respondent is interviewed, data is collected about the four preceding months. These four reference months comprise one wave. The sample in Wave 1 of the 2008 SIPP consisted of approximately 65,500 households of which only 52,029 of the households were eligible for interview. Of those eligible households, 42,030 were interviewed, with a response rate of 80.8% [18]. In subsequent waves, all adults who were interviewed in Wave 1 were followed and interviews were attempted for all household members.

Because SIPP is longitudinal, single wave response rates do not give a complete picture of nonresponse. The total sample attrition is also examined. The SIPP measures sample attrition using a sample loss rate, which is calculated by taking into account the Wave 1 nonresponse and each subsequent wave's nonresponse [18]. Given that SIPP interviews new household members joining the household after Wave 1, the growth of nonresponding households is estimated and included in the sample loss rate. The sample loss rate is 19.2% for Wave 1 and 50.5% for the final wave of the panel.

This study analyzes nonresponse bias for Waves 3 to 16 of the 2008 SIPP. The sample (i.e., households that were interviewed in Wave 1) for Waves 3 to 16 is 42,029¹ households. The analysis in this report does not include households spawned from an original sample household.

Analytic Variables

This report focuses on analysis of nonresponse bias for key estimates of the 2008 SIPP. Actual estimates of nonresponse bias can only be produced for variables that are known for both respondents and nonrespondents. Because of the longitudinal aspect of the survey, information on key estimates collected in Wave 1 can be used to analyze nonresponse bias in later waves of the survey.

The SIPP key estimates that are assessed for potential nonresponse bias include:

- Mean (monthly) household earnings
- Mean (monthly) household income
- Percent of households where at least one household member was covered by Medicaid
- Percent of households where at least one household member was covered by Medicare
- Percent of households where at least one household member received Social Security
- Percent of households where at least one household member received SSI
- Percent of households where at least one household member received SNAP benefits

In addition, variables used for household noninterview weighting adjustments are also examined [7]. These variables include:

- Age of Reference Person (Four levels: Under 25, 25-34, 35-54, 55+)
- Assets (Two levels: Bonds/Etc. at least one HH member possessed at least one of the following assets: money market deposit accounts, certificates of deposit, mutual fund shares rental property, mortgages, royalties, or other financial investments; Minimal other)
- CBSA area (Three levels: In Principal City of MSA, In MSA not in Principal City, Not in MSA or Principal city)

¹ The analysis does not include households with reference persons whose age is less than 15. The number of interviewed households including those households is 42,030.

- Education Highest level of school completed or the highest degree received by the reference person (Four levels: Less than high school; 9th-12th, no diploma; High school graduate, some college (no degree), vocational/technical school, associate degree; Bachelor's/Master's/Doctorate/Professional degree)
- Gender of Reference Person (Two levels: Male, Female)
- Household Size (Four levels: 1 persons in the household, 2 persons in the household, 3 persons in the household, 4+ persons in the household)
- Household Type (Three levels: Reference person is female with no husband present and with her own children less than 16 years old (FHHNSP), Reference person is 65 years old or older, Other)
- Income Type (Two levels: Welf/Etc. at least one household member received income from at least one of the following sources: SSI, Aid to Families with Dependent Children (AFDC), other welfare, Women, Infants, and Children Nutrition Program (WIC), SNAP, or Medicaid; Other)
- Race of Reference Person (Two levels: Black, Non-Black)
- Region (Four levels: Northeast, Midwest, South, West)
- Tenure (Two levels: Renter, Owner)
- *Urban/Rural* (Two levels: Urban area, Rural area)
- Within PSU Strata (Two levels: Low income strata, Non-low income strata) This variable was based on a probability calculated using 2000 Census Long Form income data within a designated geographic region.

Methods

The analysis involves using two different studies for measuring nonresponse bias for key estimates in Waves 3 to 16 of the 2008 SIPP. These studies utilize Wave 1 data of the 2008 Panel to examine possible nonresponse bias occurring in Waves 3 to 16 of the SIPP. The methods include comparing estimates of the full sample to the respondent sample and estimating the representivity of the survey using R-indicators.

Comparing Estimates of the Full Sample to the Respondent Sample

The first study involves comparing key estimates of the full sample (respondents and nonrespondents) to estimates of the respondent sample for each wave between Waves 3 to 16 of the SIPP. Using the variables available in Wave 1, differences between the weighted statistics (means or percentages) of the full sample and the respondent sample are examined. Initial weights, which are the weights that have incorporated unit nonresponse from Wave 1, are used in estimating the weighted statistics. Examining the differences between the estimates gives insight into the magnitude and direction of bias.

Fay's method of Balanced Repeated Replication² (BRR) is used to estimate the variance of the difference between the full sample and the respondent sample estimates. The resampling method accounts for both the complex survey design and the random error (variability) due to sampling and nonresponse.

Estimating the R-Indicator

The second analysis involves estimating the R-indicator to examine the potential for nonresponse bias. R-indicators measure how representative the respondents are compared to the original sample or population [20].

The estimation of the R-indicator involves fitting a weighted logistic regression model that predicts the probability of being a respondent as a function of Wave 1 variables to estimate response propensities. The standard deviation of the response propensities is obtained from the model, and the R indicator is estimated by the following equation [20]:

$$\widehat{R} = 1 - 2S_{\widehat{\rho}} = 1 - 2\sqrt{\frac{1}{\sum_{i=1}^{42,029} w_i - 1} \sum_{i=1}^{42,029} w_i (\widehat{\rho}_i - \overline{\widehat{\rho}})^2},$$

where

 w_i is the initial weight

 $\hat{\rho}_i$ are the response propensities estimated using the logistic regression model.

Values of the R-indicator that are close to one are an indication of strong representativeness since large values occur when the standard deviation of the response propensity is small. This means that the response propensities tend to be more similar, and therefore the respondents are more likely to be representative of the sample or population. Values close to zero are an indication of weak representativeness or that the respondents are less likely to be representative of the sample or population.

A confidence interval is constructed for the estimate of the R-indicator using Fay's method of BRR, which accounts for the sample design and the model used to estimate the response propensities [20].

² The 116 replicate weights using the Fay's method of BRR are produced in SAS using the PROC SURVEYMEANS procedure with adjustments for sample design and the VARMETHOD=BRR(FAY=0.5 OUTWEIGHTS=) statement. Initial weights are used to produce replicate weights.

Results

Table 1 summarizes the results of comparing estimates of the full sample to the respondent sample for multiple waves of the 2008 SIPP. Estimates for the full sample use the initial weight; and two estimates are calculated for the respondent sample, one weighted by the initial weight and the other weighted by the nonresponse adjusted weight. While results with the initial weight can identify the presence of nonresponse bias, the results from using the nonresponse adjusted weight will indicate if the weighting is helping to correct for some of the bias.

Relative differences are calculated for each wave comparing the current wave to the full sample estimate. The results for Waves 3, 6, 9, 12 and 15 are presented. Looking at Wave 15, 13 out of 43 estimates have a relative difference of 15% or greater. These include the following estimates: Medicare, SNAP, Social Security, West, rural, receiving welfare, owners, renters, householders < 25, householders 25-34, householders > 55, female headed household with no spouse present, and householder 65+. For some, like the householder < 25 and renters, the relative difference was greater than 15% as early as Wave 3. For other estimates, like social security and rural, the relative difference didn't reach 15% until Wave 12. Most often the nonresponse adjustment had a significant positive effect, reducing the relative difference, but the correction was generally very small.

The estimates with the largest relative differences are those for SNAP, renters, and households with reference persons less than 35. The SNAP estimate dropped from 7.6% in the full sample estimate to 5.7% for the Wave 15 respondent sample. The estimate for renter households dropped from 32% to 17% and the estimate for households with reference persons younger than 25 dropped from 4.7% to 1.3% and the estimate for households with reference persons 25-34 dropped from 15.2% to 9.1% for the Wave 15 respondent sample.

Table 2 presents results of the logistic regression model predicting the probability of a household being a respondent as a function of Wave 1 variables for Waves 3, 6, 9, 12, and 15 of the SIPP. The standard deviation of the response propensities is obtained from the model, and the R-indicator is estimated.

For Waves 3-16 of the SIPP, significant effects are observed for region, age, race, gender, education, and marital status of reference person. The South consistently has greater odds of responding than the Northeast, and the West has lower odds of responding than the Northeast across all waves. Sample units with householders who are less than 45 years of age have lower odds and units with householders who are greater than 55 years of age have higher odds of being interviewed than those who are between the ages of 45 and 54. Households with female householders have greater odds of responding than households with male householders. Compared to sample units with a White householder, units with a Black or Asian householder have lower odds of being interviewed. Households that have a householder with no high school diploma have lower odds of responding, and households with a householder with a bachelor's or a higher degree have higher odds of responding than households with a householder whose highest level of educational attainment is high school diploma or associate degree. Households with a reference person who is widowed, divorced, separated, or never married have lower odds of responding than households with a married householder.

Table 3 displays the R-indicators for Wave 3-16 of the SIPP. The R-indicator is largest at Wave 3 and smallest at Wave 12.

Table 3. R-indicators for Waves 3-16.

Wave	R-Indicator	90% CI
Wave 3	0.845	(0.8435, 0.8466)
Wave 4	0.782	(0.7799, 0.7842)
Wave 5	0.7592	(0.757, 0.7615)
Wave 6	0.7447	(0.7423, 0.747)
Wave 7	0.7207	(0.7184, 0.7229)
Wave 8	0.718	(0.7158, 0.7203)
Wave 9	0.7155	(0.7133, 0.7177)
Wave 10	0.7096	(0.7075, 0.7116)
Wave 11	0.7134	(0.7113, 0.7154)
Wave 12	0.7059	(0.7039, 0.7079)
Wave 13	0.7108	(0.7089, 0.7127)
Wave 14	0.7128	(0.711, 0.7147)
Wave 15	0.7151	(0.7133, 0.7168)
Wave 16	0.7835	(0.7823, 0.7848)

Conclusion

This analysis used two methods to examine the potential for nonresponse bias, including comparison of full sample and respondent sample estimates and estimating R-indicators. Since each method has weaknesses, no single method is sufficient on its own to draw accurate conclusions. By incorporating multiple approaches in the analysis, results were compared across methods, leading to clearer and stronger conclusions.

Some areas of potential bias were identified in the results. The relative differences between the full sample and respondent sample estimates were smallest at Wave 3 and largest at Wave 16 indicating that there may be nonresponse bias in the later waves of the SIPP. Based on the results presented, the SIPP may underestimate SNAP participation in the later waves of the panel. Other estimates potentially affected by nonresponse bias include those for households with reference persons less than 35 years of age and renters, where the later wave estimate dropped to about half of the Wave 1 estimate.

The logistic regression models that were fit to determine the R-indicators pointed to similar areas of potential bias due to certain subgroups' likelihood of responding. None of the R-Indicators for Waves 3 to 16 fell below .70 suggesting that there is a high likelihood the respondent sample is representative of the full sample or population.

A limitation of this research is that the majority of nonresponse for the SIPP happens in the first wave. Therefore, the examination of characteristics of respondents and nonrespondents in later waves may not give accurate insights into the major component of the nonresponse bias for the key estimates. In addition, results from the R-indicators are highly dependent on model assumptions and parameterization.

Many of the significant differences in the comparison of estimates are likely due to the large sample size of the survey. In classical statistics, large sample sizes lead to an increase in the probability of rejecting the null hypothesis. In the earlier waves of the SIPP, the small relative differences and large R-indicators reveal that the significant differences observed are not a cause for concern and do not directly imply that nonresponse bias exists in the earlier waves of the SIPP. However, the large relative differences and small R-Indicators in the later waves suggest that there may be the potential for nonresponse bias in the later waves of the SIPP.

References

- 1. Dahlhamer, J.M. & Simile, C.M. (2009). Subunit Nonresponse in the National Health Interview Survey (NHIS): An Exploration Using Paradata. Paper presented at *Proceedings of the Government Statistics Section of the American Statistical Association*.
- 2. Efron, B. (1982). "The jackknife, the bootstrap, and other resampling plans". Society of Industrial and Applied Mathematics CBMS-NSF Monographs, 38.
- 3. Groves, R.M. (2006). Nonresponse rates and nonresponse bias in household surveys. *Public Opinion Quarterly*, 70(5), 646–675.
- 4. Mack, S. & Petroni, R. (1994). Overview of SIPP Nonresponse Research. Paper presented at the *Proceedings for Fifth International Workshop on Household Survey Non-response*, September 26-28.
- 5. Memorandum from Ruth Ann Killion to Jason Fields. "Nonresponse Bias Analysis for Wave 1 of the 2008 Survey of Income and Program Participation (SIPP) (ALYS-12)", June 03, 2013.
- 6. Memorandum from Ruth Ann Killion to Jason Fields. "Nonresponse Bias Analysis for Wave 2 of the 2008 Survey of Income and Program Participation (SIPP) (ALYS-13)", September 02, 2014.
- 7. Memorandum from Alan R. Tupek to Chester E. Bowie. "SIPP 2004: Cross-Sectional Weighting Specifications for the Second and Subsequent Waves (WGT-5)", May 27, 2004.
- 8. Office of Management and Budget Standards and Guidelines for Statistical Surveys. http://www.whitehouse.gov/sites/default/files/omb/inforeg/statpolicy/standards_stat_surveys.pdf.
- 9. Presentation of "Practical Tools for Nonresponse Bias Studies". JPSM Short Course presented by Jill M. Montaquila and Kristen M. Olson on September 30, 2011.
- 10. Sae-Ung, S., Sissel, C. D., & Mattingly, T. L. (2007). Analytical Comparison of the SIPP and CPS_ASEC Key Longitudinal Estimates. Paper presented at the *Federal Committee on Statistical Methodology Research Conference*, November 5-7.
- 11. Sawyer A. & Perer P. (1983). The Significance of Statistical Significance Tests in Marketing Research. *Journal of Marketing Research*. Vol. XX, pp. 122 133
- 12. Schouten B., Cobben F., & Bethlehem J. (2009). Indicators for the representativeness of survey response. *Survey Methodology*, 35,101-113.
- 13. Survey of Income and Program Participation Quality Profile (1998). Available at http://www.census.gov/sipp/workpapr/wp230.pdf.
- 14. Survey of Income and Program Participation Users' Guide. Available at http://www.census.gov/programs-surveys/sipp/methodology/users-guide.html.
- 15. U.S. Census Bureau Memorandum to Nonresponse Oversight Team on March 21, 1995 from Nonresponse Research Team. Preliminary Final Report on SIPP Nonresponse.
- 16. U.S. Census Bureau Statistical Quality Standards http://cww.census.gov/msdir/docs/Quality_Standards_cww.pdf.
- 17. U.S. Census Bureau Memorandum (2002), SIPP 2000 Redesign: Creating Variance Codes for Public Use Files.
- 18. U.S. Census Bureau Memorandum (2012), Survey of Income and Program Participation (SIPP) 2008 Panel: Source and Accuracy Statement for Wave 1 to Wave 11 (core) Public Use Files. Available at http://www.census.gov/content/dam/Census/programs-

- surveys/sipp/tech-documentation/source-accuracy-statements/2008/SIPP%202008%20Panel%20Wave%2005%20-%20Core%20Source%20and%20Accuracy%20Statements.pdf.
- 19. van de Kerckhove, W., Krenze T., & Mohadjer L. (2006). Approaches to a Nonrespones Bias Analysis in an Adult Literacy Survey. *Proceedings of the Survey Research Methods Section, American Statistical Association*, 3790-3795.
- 20. Witt, M. B. (2010). Estimating the R-indicator, Its Standard Error and Other Related Statistics with SAS and SUDAAN. Paper presented at *JSM Proceedings, Section on Survey Research Methods. American Statistical Association*.
- 21. Young, P. (2010). "Jackknife and Bootstrap resampling methods in statistical analysis to correct for bias," lecture notes, Physics Department, University of California Santa Cruz, Santa Cruz, CA.

۵
2008 SIPP
8
8
e 2
£
ð
e
Ē
Sa
t
de
6
Sp
Re
P
a
픙
Ē
Š
3
F
ō
sf
ite
Ë
sti
fE
0
S
ari
ğ
ó
_:
e 1
qe
۳

lable 1. Cor	Table 1. Comparison of Estimates from Full Sample and Respondent Sample of the 2008 SIPP	rilmates Ti	Tom ruii sampie and Res	aldm	and kespo		Wante December	or the	e 2008 SIPP		Standand O overW	100	pope		standand Ct our/M	buons	4	Mario	15 00	Many of Bornand		
	ridings in		Initial		Nonresp	80	Initial	_	Nonresp	11	Initial	_	Nonresp		Initial	Non	Nonresp	Initial		Nonresn	, 5	
	Initial Weight	eight	Wgt		Adj Wgt		Wgt	- 0:-	Adj Wgt		Wgt	. ~	Adj Wgt		Wgt	Adj Wgt	Wgt	Wgt		Adj Wgt	يب ١	
			Rel		Rel		Rel		Rel		Rel		Rel		Rel	Rel	<u> </u>	Rel		Rel		
:	Percent/	Std	Diff		Diff		Diff		Diff	-	Diff		Diff		Diff	Δ :	Diff	Diff		Diff		
Variables	Mean	Error	8	+	(%)	+	(%)	+	(%)	+	<u>%</u>	+	<u>%</u>	+	(%)	%	(9)	8		8		
Medicare	27.82	0.25	6.04	*	3.49	*	11.91	*	7.24	*	14.87	*	8.55	*	* 68.91	9.26	* 97	17.82	*	9.07	*	
Medicaid	17.17	0.22	-4.40	*	-2.27	*	-7.94	*	-6.80	*	-10.84	*	* 16.8-	*	-13.68 *	-11.62	* 79.	-14.37	*	-12.63	*	
SNAP	7.56	0.14	-8.56	*	-6.28	*	-15.78	*	-14.98	*	-24.80	*	* -23.15	*	+ 27.97 *	-27.18	* * *	-33.47	*	-32.58	*	
SSI	4.57	0.11	-0.72		0.58		-0.34		-0.82		-2.85		-3.49	-	-2.66	ώ.	-3.95	-3.78		-5.47	*	
Social Security	30.14	0.26	5.42	*	3.18	*	11.07	*	6.91	*	13.79	*	8.07	*	15.99 *	8.96	* 96	16.91	*	8.77	*	
										-												
Household Earnings	\$4,307.26	39.81	1.47	*	1.66	*	1.44	*	3.80	*	1.81	*	4.99	*	3.12 *		* 80.7	2.61	*	7.13	*	
Household Income	\$5,391.53	41.07	2.06	*	1.84	*	3.19	*	4.45	*	4.31	*	5.95	*	5.70 *	7.70	* 02	6.04	*	8.20	*	
																						
Region												- 55										
Northeast	18.85	0.17	0.89		1.31	*	2.72	*	3.45	*	5.66	*	3.84	*	* 69.8	4.62	52 *	3.66	*	4.92	*	
Midwest	22.9	0.24	1.95	*	1.06	*	0.89		98.0		0.41		0.03		1.39	1.11	11	1.08		1.00		
South	35.94	0.27	1.05	*	96.0	*	3.50	*	2.97	*	4.89	*	4.44	*	* 06.4	4.34	* *	5.32	*	4.62	*	
West	22.32	0.22	-4.76	*	-3.98	*	-10.10	*	-9.77	*	-12.43	*	-12.28	*	* -14.97	-14.41	.41 *	-15.52	*	-15.27	*	
Irhan/Rura																						
Rural	20.6	0.58	r X	*	7 60	*	0 10	*	27.8	*	12.29	*	11 98	- A-	* * * * * * * * * * * * * * * * * * * *	14 56	*	16 19	*	15.91	*	
Urban	79.4	0.58	-1.56	*	-1.27	*	-2.70	*	-2.55	*	-3.77	*	* 99.5-	ж.	* 77.4-		£ *	-5.28	*	-5.16	*	
Within PSU Strata																						
Non-low income Strata	76.8	0.32	1.13	*	0.73	*	1.59	*	1.58	*	2.25	*	* 2.26	*	2.55 *	2.54	*	2.74	*	2.82	*	
Low income Strata	23.2	0.32	-3.93	*	-2.48	*	-5.64	*	-5.61	*	-8.24	*	* -8.27	*	.9.48 *	-9.44	* 44	-10.29	*	-10.62	*	
1																						2.50
<u></u>		-		-88		=0		- 10		=		-		=				-		- 155 - 155		200

(Table 1 continued)	All Sample Cases	Cases	Wave 3 Respondents	Respo	undents		Wave 6 Respondents	espon	dents		Wave 9 Respondents	espoi	dents		Wave 12 Respondents	Respo	ndents	_	Wave 15 Respondents	Resp	ondents	
	Initial Weight	eight	Initial Wgt		Nonresp Adj Wgt		Initial Wgt	Z 4	Nonresp Adj Wgt		Initial Wgt	2 4	Nonresp Adj Wgt		Initial Wgt	2 4	Nonresp Adj Wgt		Initial Wgt		Nonresp Adj Wgt	
Variables	Percent/	Std	Rel Diff (%)		Rel Diff (%)		Rel Diff (%)		Rel Diff (%)		Rel Diff (%)		Rel Diff (%)		Rel Diff (%)	<u> </u>	Rel Diff (%)		Rel Diff (%)		Rel Diff (%)	
CBSA Type						 				 				\vdash					-			
In Principal City of MSA	32.83	0.75	-4.21	*	-3.17	*	-7.14	*	-6.90	*	-8.85	*	-8.68	*	-11.53	*	-11.32	*	-12.35	*	-12.11	
In MSA not in Principal City	50.3	0.92	1.39	*	1.06	*	2.65	*	2.82	*	2.66	*	2.83	*	3.41	*	3.60	*	4.01	*	4.20	
Not in MSA or principal City	16.86	1.29	3.60	*	2.77	*	4.68	*	3.81	*	7.18	*	6.48	*	8.81	*	8.02	*	8.25	*	7.42	
HH size								550		*												
1	28.63	0.27	-2.31	*	-2.85	*	. 69.8-	*	-7.10	*	-5.84	*	-10.63	*	-8.16	*	-14.17	*	-10.70	*	-17.19	
2	34.74	0.28	2.14	*	1.54	*	74	*	3.06	*	5.46	*	4.53	*	6.74	*	5.28	*	8.15	*	6.44	
3	15.08	0.17	-0.97	*	0.11		-2.06	*	1.21		-2.37	*	1.27		-1.08		3.49	*	-0.52		3.71	
4+	21.54	0.2	0.19	******	1.12	*	-0.08		2.82	*	-0.31		4.09	*	-0.85		4.78	*	-1.07		5.44	
Assets																						
Bonds/Etc.	39.05	0.32	4.57	*	3.51	*	7.68	*	7.58	*	10.16	*	9.76	*	11.53	*	11.26	*	12.84	*	12.53	
Minimal	60.95	0.32	-3.17	*	-2.39	*	5.63	*	-5.55	*	-7.81	*	-7.45	*	-9.11	*	-8.85	*	-10.42	*	-10.10	
Income Type				100 - 11																		
Welf/Etc.	18.98	0.24	-4.80	*	-2.72	*	. 8.45	*	-7.40	*	-11.59	*	-9.83	*	-14.48	*	-12.55	*	-15.54	*	-13.82	
Other	81.02	0.24	1.06	*	0.62	*	1.79	*	1.59	*	2.37	*	2.05	*	2.88	*	2.55	*	3.05	*	2.77	
Tenure																- 477						
Owner	68.02	0.28	6.30	*	4.95	*	11.91	*	11.17	*	14.86	*	14.02	*	16.79	*	16.05	*	18.19	*	17.48	
Renter	31.98	0.28	-16.67	*	-12.47	*	-40.35	*	-36.49	*	-59.03	*	-53.08	*	-75.21	*	-68.52	*	-89.77	*	-81.96	

(Table 1 continued)	All Sample Cases	ases	Wave 3 Respondents	sponde	nts	Wav	e 6 Res	Wave 6 Respondents		Wave 9 Respondents	puods	ents	- M	ive 12 R	Wave 12 Respondents	ints	3	Wave 15 Respondents	espon	dents	
			Initial	Nonresp	esb	Initia	_	Nonresp		Initial	2	Nonresp	=	Initial	Nonresp	resp		Initial	ž	Nonresp	
	Initial Weight	ght	Wgt	Adj Wgt	Vgt	Wgt		Adj Wgt		Wgt	Ad	Adj Wgt	5	Wgt	Adj Wgt	Ngt	_	Wgt	A	Adj Wgt	
			Rel	Rel	_	Rel		Rel		Rel	_	Rel	<u> </u>	Rel	Rel	<u>-</u>		Rel		Rel	
	_	Std	Diff.	Diff	:	Diff		Diff		Diff		Jiff S	Δ :	Diff	Diff	± ։	_	Diff		Diff	
Variables	Mean	Error	(%)	%		8		<u>%</u>	1	(%)	1	(%)		(%)	%		1	(%)	-	(%)	T
Age of Householder									_												-
Under 25	4.69	0.13	-34.23 *	-31.44	* 44	-112.82	32 *	-111.99	*	-170.18	* -16	-162.17 *		. 736.90	* -225	-225.49	* -2(* 96.092-	*	-241.64	*
25-34	15.21	0.21	-12.03 *	-9.80	* 00	-28.83	*	-24.87	*	.43.07	<u>۴</u>	* -37.31	-5.	-54.15	* -46.66	99	۹ *	* +66.74	*	-57.52	*
35-54	39.97	0.25	* 0.84	1.86	* 9	1.11	*	3.62	*	66.0	*	4.63 *	о́ —	0.45	2.60		*	97.0	10.005	7.07	*
55+	40.13	0.27	* 2.84		* 2	11.94	*	80.6	*	, 60.21	*	11.27 *		17.41	* 12.47		*	* 09.81	*	12.69	*
7.001								- 100 01													
Gender of Householder								-													
Male	46.82	0.25	-0.28	-0.39	62	-0.18	~	-0.11		0.32	-	0.48	0	0.62	0.87	37	_	0.37		0.65	
Female	53.18	0.25	0.24	0.34	4	0.16		0.10		-0.28	<u> </u>	-0.42	٩	-0.55	-0.78	78	<u>ү</u> —	-0.32		-0.58	
Race of Householder																					
Non-Black	88.09	0.23	0.62 *	0.18	*	0.64	*	95.0	*	1.08	*	* 66.0		1.22	* 1.10		*	1.18		1.18	-
Black	11.91	0.23	* +84	-1.35	* 52	-5.01	*	-4.31	*	* 97.8-	*	* 96.7-	-	* 80.01-	* -8.92		*	-9.73		-9.70	
НН Туре									-												
FHHNSP	1.22	80.0	* -8.54	-6.68	*	-16.90	*	-13.41	*	-29.87 *		-23.52 *		-29.40	* -19.75		? *	-27.71 *		-18.32	*
Reference person is 65 years old or older	22.23	0.23	6.53 *	3.11	*	12.95	*	66.9		16.05 *		7.82 *		17.88	* 7.89		*	* 18.29		6.71	*
Other	76.55	0.25	* *************************************	-0.84	*	-4.26	*	-2.04	* *	* -5.47		-2.21 *		-6.34	* -2.28		Ψ *	* -6.56		-1.88	*
																					25
											<u> </u>										
						_			_				_				_				

(Table 1 continued)	All Sample Cases	Cases	Wave 3 Respondents	Respo	undents	F	Wave 6 Respondents	noas	dents		Wave 9 Respondents	espor	dents		Wave 12 Respondents	espoi	ndents		Wave 15 Respondents	lespon	dents	
			Initial		Nonresp	_	Initial	. ž	Nonresp		Initial	.—	Nonresp		Initial		Nonresp		Initial	Ž	Nonresp	
	Initial Weight	eight	Wgt		Adj Wgt		Wgt	Ă	Adj Wgt		Wgt	4	Adj Wgt		Wgt	4	Adj Wgt		Wgt	AG	Adj Wgt	Т
			Rel		Rel		Rel		Rel		Rel		Rel		Rei		Rel		Rel		Rel	_
	Percent/	Std	Diff		Diff		Diff		Diff		Diff		Diff		Diff		Diff		Diff		Diff	
Variables	Mean	Error	(%)		(%)		(%)		(%)		(%)		(%)	+	(%)	-	(%)		(%)		(%)	\neg
Educational Attainment of Householder	Householder																					
Less than High School	4.61	0.14	0.88		1.52		5.19	*	2.90	*	5.65	*	2.97	*	6.72	*	2.89	88.0	7.90	*	3.42	
High School, no diploma	6.75	0.14	-5.22	*	* 07.4-	*	* -5.27	*	-7.58	*	-6.14	*	-8.14	*	* 27.72	*	* 60.01-	*	-10.72	*	-14.42	*
High School graduate	58.64	0.32	-0.32		-0.35		* -1.03	*	-1.09	*	-1.48	*	-1.46	*	-1.75	*	-1.86	*	-1.49	*	-1.59	*
College graduate	30	0.32	1.58	*	1.43	4	2.23	*	3.13	*	3.13	*	3.87	*	3.73	*	4.91	*	3.59	*	5.08	*
0	trainitaed meraporal has semested to receive the second second of the second se		- concord je	O Pur	itred merro.	+	0000 40	Danel	Ear info	rmatio	mes do d	nling	and page	mulin	tion 2009 Danal Ear information on campling and population perfor see	а						

Source: U.S. Census Bureau, Survey of Income and Program Participation, 2008 Panel. For information on sampling and nonsampling error see

http://www.census.gov/programs-surveys/sipp/tech-documentation/source-accuracy-statements.html.
* The estimate for respondents is significantly different from the full sample estimate at the 90% confidence level.

Table 2. Odds Ratios from Logistic Regression Model

Wave	5	Wave 3	8	Wave 6	8	Wave 9	3	Wave 12	Wa	Wave 15
Model Variables	Odds Ratio	ID %06	Odds Ratio	12 %06	Odds Ratio	12 %06	Odds Ratio	12 %06	Odds Ratio	90% CI
Region										
Northeast†	1	Ī	T	Į.	Н	1	Н	1	Н	I
Midwest	1.11*	(1.01, 1.21)	0.98	(0.91, 1.06)	0.99	(0.92, 1.06)	0.99	(0.93, 1.06)	0.99	(0.93, 1.05)
South	1.06	(0.99, 1.14)	1.07*	(1.01, 1.14)	1.12*	(1.06, 1.18)	1.08*	(1.02, 1.14)	1.08*	(1.02, 1.14)
West	0.79*	(0.73, 0.86)	0.73*	(0.69, 0.78)	0.75*	(0.7, 0.8)	0.73*	(0.69, 0.78)	0.75*	(0.7, 0.8)
Age		TOTAL TOTAL STREET, AND STREET, STREET		Andrews Comp. Structure Company		100 mm anno 100 mm			1	
15-24	0.31*	(0.28, 0.35)	0.21*	(0.19, 0.23)	0.18*	(0.16, 0.2)	0.16*	(0.15, 0.18)	0.16*	(0.14, 0.18)
25-34	0.5*	(0.46, 0.54)	0.45*	(0.42, 0.47)	0.41*	(0.38, 0.43)	0.4*	(0.38, 0.42)	0.37*	(0.35, 0.39)
35-44	0.73*	(0.68, 0.8)	0.73*	(0.69, 0.77)	*69.0	(0.66, 0.73)	0.7*	(0.66, 0.74)	*4.0	(0.64, 0.7)
45-54†	н	1	1	1	н	1	Н	1	1	;
55-64	1.12*	(1.03, 1.22)	1.23*	(1.16, 1.3)	1.24*	(1.16, 1.32)	1.3*	(1.22, 1.38)	1.28*	(1.21, 1.35)
+59	1.34*	(1.24, 1.46)	1.44*	(1.35, 1.53)	1.44*	(1.35, 1.53)	1.47*	(1.38, 1.56)	1.37*	(1.3, 1.45)
Gender										
Male†	1	1	1	-1	1	1	Н	1	н	1
Female	1.08*	(1.04, 1.13)	1.06*	(1.02, 1.1)	1.05*	(1.01, 1.08)	1.05*	(1.01, 1.1)	1.07*	(1.03, 1.11)
							-			
Race										
White†	1	1	1	1	1	1	П	1	П	ı
Black	0.81*	(0.75, 0.87)	0.92*	(0.86, 0.98)	0.87*	(0.82, 0.93)	.88*	(0.83, 0.94)	0.91*	(0.86, 0.97)
Asian	0.77*	(0.69, 0.87)	0.79*	(0.72, 0.87)	0.83*	(0.76, 0.92)	0.82*	(0.74, 0.91)	0.87*	(0.78, 0.97)
Other	0.83*	(0.74, 0.93)	*6.0	(0.81, 1)	0.85*	(0.77, 0.94)	0.83*	(0.74, 0.93)	0.84*	(0.74, 0.95)
Hispanic Origin										
Hispanic origin†	F	1	Ţ	1	1	1	Н	ļ	7	ı
Non-Hispanic origin	Н	(0.92, 1.08)	0.92*	(0.86, 0.98)	0.85*	(0.81, 0.89)	0.88*	(0.83, 0.93)	0.85*	(0.8, 0.91)
Education										
Codos drid codes	30.0	10 96 1 07	1 05	(0.07.1.15)	•	(0 0 1 00)	2	(0.95 1.14)	20	(0.96.1.14)
0+h 12+h mode No dialoma	*05.0	(0.00, 1.07)	T:00	(0.97, 1.13)	***************************************	(0.32, 1.05)	*000	(0.93, 0.95)	*500	(0.30, 1:17)
otil-tztii glade, ivo dipioliia	67.0	(0.73, 0.03)	0.0	(0.01, 0.34)	0.00	(0.00, 0.00)	0.0	(000, 000)	9	(20:0) (0:0)
High school graduate/Associate degree†	1	ľ	1	l	н	E	П	Ē	Т	1
Bachelor's or higher degree	1.11*	(1.05, 1.18)	1.11*	(1.06, 1.17)	1.13*	(1.08, 1.18)	1.13*	(1.08, 1.17)	1.1*	(1.05, 1.15)

(Table 2 continued)	M	Wave 3	3	Wave 6	3	Wave 9	W	Wave 12	Wa	Wave 15
Model Variables	Odds	12 %06	Odds	12 %06	Odds	12 %06	Odds Ratio	ID %06	Odds Ratio	ID %06
Marital Status										
Married [†]	Т	1	1	1	Н	;	Н	1	Н	1
Widowed	0.73*	(0.66, 0.8)	0.7*	(0.65, 0.76)	*99.0	(0.61, 0.71)	.59*	(0.55, 0.63)	0.59*	(0.55, 0.63)
Divorced	0.7*	(0.66, 0.75)	*69.0	(0.65, 0.73)	*19.0	(0.63, 0.7)	.65*	(0.62, 0.68)	0.64*	(0.61, 0.68)
Separated	0.57*	(0.51, 0.64)	0.53*	(0.47, 0.6)	0.51*	(0.47, 0.56)	0.46*	(0.42, 0.5)	0.47*	(0.42, 0.53)
Never married	0.71*	(0.67, 0.75)	0.7*	(0.67, 0.73)	*89.0	(0.65, 0.72)	*49.0	(0.64, 0.71)	.68*	(0.65, 0.72)

Source: U.S. Census Bureau, Survey of Income and Program Participation, 2008 Panel. † Reference category * P-value < 0.10